Quantifying Side-Channel Information Leakage from Web Applications

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Abstract. Recent research has shown that many popular web applications are vulnerable to side-channel attacks on encrypted streams of network data produced by the interaction of a user with an application. As a result, private user data is susceptible to being recovered by a side-channel adversary. A recent focus has been on the development of tools for the detection and quantification of side-channel information leaks from such web applications. In this work we describe a model for these web applications, analyse the effectiveness of previous approaches for the quantification of information leaks, and describe a robust, effective and generically applicable metric based on a statistical estimation of the mutual information between the user inputs made in the application and subsequent observable side-channel information. We use our proposed metric to construct a test capable of analysing sampled traces of packets to detect information leaks, and demonstrate the application of our test on a real-world web application.

1 Introduction

A pervasive theme in recent web application development is a move toward increasingly complex, interactive programs, with more and more computation being moved into the cloud. This has resulted in an increase in user data being retained in cloud services and accessed by the user using a web browser on any type of client device. Consequently, a significant amount of personal information is now being transferred across networks as users interact with these web applications day-to-day.

A passive adversary eavesdropping on the network traffic between a client and the web application server will be able to capture IP packets and reconstruct the data being sent to and from the web application if the data is unencrypted. If the web application is processing user-entered data that a user wishes to remain private, the server hosting the web application will typically deploy the TLS protocol over HTTPS to encrypt the data with the goal of ensuring that an adversary eavesdropping on the traffic resulting from web application transitions cannot learn the inputs entered by the user or the particular state the web application is in.

However, as first pointed out by Chen et al. [8], side-channel observations made on network traffic between a client and a server during the execution of a web application can reveal information about the user inputs made at a particular state or personal information stored in server databases. Chen et al. illustrated that an adversary can utilise knowledge of tuples of packet sizes and directions to reveal user choices made in financial, medical and search applications.

Follow-on work has focused on developing strategies and tools for web application designers and implementors with which to find such information leakages in order to secure their applications against side-channel adversaries [5,32]. A key point in this work is to spot and quantify when/where the information leaks from a web application within the network traffic. The approaches currently available all come with several caveats, such as high dependency on chosen metric [5] or the inability to deal with noisy observations [32].

Our work is motivated by the desire to avoid the aforementioned caveats. Our proposed solution is to model the system as a classical information-theoretic channel, with the application inputs considered as the hidden channel inputs and any data available to an eavesdropping network adversary as the channel output. We use statistical sampling to produce an estimate (and, in the case of discrete data, confidence intervals) for the mutual information between the web application inputs and side-channel observations. One advantage of using mutual information over alternative heuristic-based classifiability metrics is that it allows us to compute the information leakage for a given usage pattern for the web application, or by computing the channel *capacity* allows us to determine the usage pattern of the application that leaks the most information. A second advantage of our mutual information metric is that in the discrete case derived confidence intervals provide a degree of rigour unavailable in previous approaches to solving this problem.

We outline a suitable leak-detection technique capable of mechanically and quickly identifying the majority of simple information leaks. Our technique involves considering the time-series of the sequence of packets produced by the communication between the client computer and the web server and computing the mutual information between the application inputs and selected properties of the packets at each position in the sequence. Considering the distributions of individual packets allows us to quickly detect the most common types of information leak; the joint distribution of multiple packets can be computed to detect more complex leaks.

This paper is structured as follows. We provide a formal model for a web application in section 2 and review related literature. In section 3 we discuss previous approaches and related work for this problem. Our choice of metric is detailed (as well as its pros and cons) in section 4. Thereafter we formally describe our leakage detection test and illustrate it on (a part of) the NHS direct website¹ in section 5. Finally, in section 6 we demonstrate further the application of our test in analysing the effectiveness of several countermeasures.

2 Modelling web applications

Every web application can be formalised as an input-driven finite state machine. A web application always exists in some state that waits for some user-specific input that may be one or more combinations of mouseclicks, form elements, database data or key presses. The combination of these inputs determines the new state the web application should transition to. Inputs may be selected on both the client and server machines, either directly by the user (for example through selecting hyperlinks or filling out HTML forms), generated by the web application itself (e.g. cookies, unique identifer strings or internal program variables), or taken from database storage (e.g. user account information). We hence formally describe a web application as follows below.

Model. A web application may be represented as the following quintuple $(\mathcal{S}, \mathcal{O}, \mathcal{A}, f, g)$:

- -S A web application is always in one of m = |S| possible states $s \in S$. We define a state to be a unique configuration of the web browser DOM on a client machine.
- \mathcal{A} Each state is associated with a set of inputs \mathcal{A} . At each state a subset of inputs $a \in \mathcal{A}$ are sent to server and used to determine the next state the web application should transition to.
- -f Each transition from one state to another is defined according to some (potentially probabilistic) function $f: S \times A \rightarrow S$. The evaluation of f is made by the application web server, and the next application state is transferred back across the network to the client machine.
- $-\mathcal{O}$ State transitions cannot be observed directly, but each transition is associated with one of $n = |\mathcal{O}|$ observable side-channel outputs $o \in \mathcal{O}$. We call these outputs a **trace**. Note that each observable output is a single trace.
- -g The observable outputs are generated according to a function $g: S \times A \to O$, where $g(s, a) = o \circ \epsilon$, with ϵ representing the noise in the observations.

We now discuss further the properties of our observables, i.e. traces, in practice.

¹ This is a free web application provided by the UK national health service. The part of the application that we investigate is a confidential symptoms checker.

2.1 Traces

Traces are the result of a state transition within a web application: when an input to a particular web application state is made by a user, the client's browser first sends the input data to the server hosting the web application, which then responds with additional data or a new web page corresponding to the next web application state. The nature of this outgoing and incoming traffic to the client is therefore dependent on the particular data being entered into the web application and the corresponding state the web application is in.

Non-deterministic traces. A key factor to consider is the non-deterministic nature of network traces generated from the selection of an input at a particular state. Variation in sampled traces comes from the fact that in practice, network traffic between a client machine and a server hosting a web application is dependent on many factors. For instance:

- Network quality/congestion/latency and the user's geographic location at the time of data capture will
 affect total traffic size and packet arrival timings;
- Web pages themselves vary over time, for instance through dynamic content, different advertisements and varying time-dependent information, which affects total traffic size;
- The connection type and medium ('normal', VPN, Tor; fibre-optic, satellite) will heavily affect all types of packet features;
- Traffic information may be added at layers below the application layer in the network model, which may affect packet sizes;
- Users may click to transition to a new state before the current state has finished loading, artificially shortening the trace;
- Proposed application, transport and link-layer countermeasures typically attempt to randomly distort network traces.

Clearly it becomes important to either control for these factors within any leakage detection strategy or ensure that they are clearly defined as assumptions in the implementation configuration, e.g. : *This web application has leakage X when a user connects through Tor.* The division of the factors into these two categories is determined by whether a factor can be accounted for in the description of the implementation environment or not. For example, variation in web page content as a result of advertising is difficult to control for, but will result in variable trace sizes when the web application is repeatedly executed with the same inputs. Summarising: it is critical that any leakage detection strategy can incorporate probabilistic output traces.

Trace features A network trace *o* is a multi-dimensional object, consisting of multiple packets each of which have multiple properties: for example a hypothetical trace may consist of 20 TCP packets, each of which have their own arrival time, payload size, direction (to the server, to the client) and set TCP options. We define a trace *feature* to be a particular function of a subset of packet properties taken over all the packets in a trace. Example features may be the total size of all of the trace packets, the size of a particular packet, the order of the communication between the client and the server or inter-packet arrival timings. Figure 1 illustrates a trace sampled from our online health website.

Not all of the information contained within a network trace will be dependent on the input choices made by the user. In many cases the inter-packet arrival timings may be entirely independent of the input choice made by a user, for instance. Taking the perspective of a Bayesian adversary, the challenge is to identify the trace features that most strongly distinguish one input choice from another, and then use these efficiently in an attack algorithm. Dyer et al. make a similar claim in their recent work on subverting common traffic analysis countermeasures [12]. From the perspective of a person attempting to assess the information leakage of the application, the challenge is somewhat different: to identify the trace features that leak the most information about the input choices and if possible mitigate this leakage.

Example. As a toy example, we will consider a very simple hypothetical web application consisting of three states S_1 , S_2 and S_3 . At S_1 the user picks one of two inputs a and ab. If the user chooses a, then the web



Fig. 1. A packet trace sampled from a popular online health website. The total size in bytes of each individual packet is represented as the trace feature.

application transitions to the new state S_2 , and in the case of b transitions to S_3 . A key difference between the two states S_2 and S_3 is that S_3 is approximately 10000 bytes in size, 2000 bytes larger than S_2 at 8000 bytes. In this case, trace features that will leak information are the total size of the server response, and the presence of additional packets in the longer traces produced when the user selects b. A visual representation of the application is in Figure 2.



Fig. 2. A simple hypothetical application with three states, two inputs and differing server response sizes.

2.2 Designer and adversary strategies

The steps required in the processes of quantifying information leakage from a designer's perspective and attacking a web application from an adversarial perspective are outlined in Figure 3. We can see that the first two steps are equivalent for the roles of designer and adversary. As in practice a web application has to be treated like a black-box, to understand the behaviour of the application outputs both roles have to *sample* the web application in its deployed state. It is critical for both roles to sample the application well enough to accurately estimate the trace behaviour. A different way of describing this is to say that the success of an attack or leak quantification method is essentially dependent on the underlying ability of the designer or adversary to successfully estimate the probability distribution of the trace outputs given the secret inputs. An adversary may or may not have full access to every area of the application, so in some cases the adversary's ability may be diminished in this area. As network traces are significantly multi-dimensional, time-ordered vectors of features, the next challenge for an adversary and a designer is to identify the feature(s) or transformation of features that contain the most information about the secret inputs. This stage is highly similar to the feature extraction and feature selection problems in machine-learning.

The final stage for the designer is to quantify the information contained in the selected trace features. The adversary is faced with a *classification* problem; given new data, he or she has to guess the class of inputs



Fig. 3. Schematic defining the differing approaches taken by an application designer looking to identify information leaks, and an adversary attempting to attack the application.

the data resulted from. In this supervised learning context, empirical studies indicate that the most common attack strategy is a Bayesian classifier [3,16,18,23], so typically the adversary next has to select an appropriate classification strategy, and train the resulting method using the data sampled in step 1. Finally, the adversary will execute their attack using one or more traces captured during the interaction of a user with the web application.

So whilst an application designer attempting to detect information leaks and an adversary attempting to attack an application have the dual goal of identifying (and thus quantifying) the trace features that leak the most information, the feedback required by the designer and the adversary is different. An adversary is concerned with the power of his or her attack, and hence requires a metric capable of quantifying the practical performance of the attack strategy. The designer is concerned with the size of the information leak itself; not with precisely evaluating the exact ability of any single adversary to exploit it.

3 Previous work

Previous work in the area focuses on two distinct questions. The first is how to construct successful attacks (i.e. the focus is on exploitation of leakage; we described this as the attacker's perspective previously). The second question is how to detect and mitigate leakage (i.e. the focus is on leakage detection; we described this as the designer's perspective previously).

It is critical that any leakage metric be sufficiently generic to be able to directly compare two different implementations of an application. The performance of attack strategies will vary across different applications, and hence by recording the success of an attack alone we can only make comparisons that assess information leakage against that particular adversary. The most common attack strategy is a Bayesian classifier, and that of the various inputs to the classifier that affect its success, the choice of trace feature(s) is paramount [12]. Hence, we suggest that a suitable leakage metric should be able to quantify the information contained in a set of trace features. Standaert et al. take the same approach of distinguishing between metrics for assessing security and metrics to quantify the quality of attacks in their framework for the evaluation of side-channel key-recovery attacks [27].

We now focus on previous work concerned with leak detection and leakage metrics.

3.1 Leak detection/quantification

There have been two different approaches towards quantifying side-channel information leakage in the context of web applications: using information theory to estimate the reduction in entropy of the application inputs after the adversary has seen the observable outputs, and using metrics to quantify the classifiability of the secret inputs given the observable outputs.

Chen et al. [32] propose using the conditional entropy of the secret inputs after the trace outputs have been observed. They describe *rerun testing*, which samples the channel distribution by running the web application source code, and then computes the conditional entropy between the application inputs and the sampled observations. The conditional entropy metric estimates the remaining uncertainty about the tested inputs given the sampled observations, and so can be used directly as a quantification metric.

Rerun testing is a white-box strategy that does not incorporate any probabilistic trace information; in the data acquisition phase any trace features that vary are discarded. Noise will be present in a situation where a web application is deployed as a live system, as described in section 2, so a strategy of retaining only invariant trace features will often result in valuable information being discarded. In our own experiments we have observed situations in which strongly distinguishing packets are sometimes noisy; an adversary will only sometimes be able to determine the input choice made, but the information leak may be significant.

Chapman and Evans [5] describe two different classifiability-based quantification metrics that estimate information leakage by clustering sampled trace features using several distance metrics derived from an analysis of trace features, and then a quantify the quality of the clustering to capture the size of a leak. Informally, the hypothesis is that clusters that are more strongly separated should produce a lower estimate of information leakage.

Their first metric estimates the *ambiguity set reduction* of the input space in terms of loss of entropy. An ambiguity set is a set of inputs that an adversary cannot distinguish from each other. To decide when two inputs are indistinguishable, Chapman and Evans take the center of each cluster of sampled outputs and determine a threshold distance around each cluster center that includes a certain percentage of the sampled outputs for that input. If the center of a cluster is inside the threshold distance of another cluster, then those two clusters are considered indistinguishable to an adversary. An inherent problem with this strategy is that the leakage result becomes very sensitive to the almost arbitrary choice of threshold distance. The only feedback mechanism to aid in the establishment of the threshold distance is the success of an attack; hence it is likely that a leakage result calculated using this method will estimate the strength of the *attack*, rather than the size of the information leak itself. Additionally, the linearity of the probability distribution of the features is not taken into account. In some situations, there may be no natural ordering to the feature distribution; treating the sampled features as continuous, ordered data will in some cases not make use of this property.

Chapman and Evans' second metric directly estimates the clustering quality by computing the Fisher criterion ([14]), a function that is typically used in linear discriminant analysis (LDA) to estimate classifiability for use in feature extraction. The Fisher criterion computes the ratio between the inter- and intra-cluster variance; hence two heavily overlapping clusters will have a low Fisher criterion, and as clusters become more strongly separated, the Fisher criterion will tend to infinity. The Fisher criterion is a capable metric in the case of two clusters being linearly separated, but its suitability will degrade as the relationship between two clusters becomes increasingly non-linear, or as the sampled data becomes decreasingly less continuous. In general, it may be that a particular classifiability metric works well for a data set captured from one web application, but not another, or works well for a particular type of trace feature and not others. Any classifiability result is sensitive to the choice of distance metric for the feature. For continuous data, such as packet sizes, the choice of distance metric for the feature. For continuous data, such as packet sizes, the choice of distance metric for the features. For continuous data, such as packet sizes, the choice of distance metric metric for the features such as the control flow of the trace data, the answer is less obvious, and for trace features considered jointly, the answer becomes increasingly difficult.

A general limitation with a leakage result derived from an approach attempting to quantify classifiability is that it cannot include any information contained in the a-priori distribution of the inputs. If the users of a web application follow a particular usage pattern, the leakage metric cannot incorporate this additional information.

Example. Returning to the toy example outlined in section 2, we can construct a situation where the classifiability and entropy metrics as described will fail to detect a leak. Let us now consider a scenario whereby the size of each state is as before, but an object approximately 6000 bytes in size (perhaps a particularly large advert)



Fig. 4. (Left) Our modified toy example, with strongly varying server response sizes for each input choice. (Right) Visualisation of sampled data for our modified toy example. The samples corresponding to the different inputs are represented by light circles and squares, and the centroids are marked by the darker symbols.

appears at both s_2 and s_3 50% of the time, independently of the input choice made by the user. Therefore, the traces produced when the user selects a are approximately 8000 bytes in size half the time, and 14000 bytes in size in the rest of the time. The traces produced as a result of b are 10000 and 16000 bytes in size respectively. The left-hand side of Figure 4 gives a visual representation of this modified application. The centroid of the traces corresponding to a will be approximately 10000 bytes, and for b, 12000 bytes.

The inter-class variance now will be significantly smaller than the intra-class variance, resulting in a Fisher criterion score significantly less than 1 (in fact, assuming 200 samples and low noise, approximately 0.1). A low Fisher criterion score is difficult to interpret meaningfully, and here could quite easily be interpreted as evidence for the sets of traces being indistinguishable. However, it is clear that an adversary observing a trace size close to one of the 4 possible values can always determine whether the user chose a or b. The right-hand side of Figure 4 indicates visually using hypothetical sampled data why the linear Fisher criterion metric fails to detect a leak. The samples corresponding to input a are represented by circles, and for b by squares. With this distribution of the outputs there is no way to linearly separate the two classes of samples. We can see straightforwardly that a similar failure to detect the leak will occur using the centroid-based entropy metric.

Similar situations to our toy example can occur in practice; any fixed-size object that randomly occurs in traces has a chance of creating non-linear cluster relationships. As a more subtle example, packet re-transmissions can result in trace sizes being increased by multiples of (approximately) a fixed amount. This can result in the distribution of trace sizes being only locally continuous; any underlying difference in the sizes of two traces has a good probability of being preserved despite the presence of the re-transmissions. In this situation, clusters of trace sizes will have non-linear relationships.

4 Mutual Information based Leakage Detection

We hope to have convinced the reader that there is little hope in finding a generically applicable metric through a quantification of feature classifiability. We hence advocate the use of a different approach all together. In what follows we assume the reader knows basic facts about information theory².

The problem of inferring the input choices made by a user from within the set of all possible inputs bears a significant resemblance to the problem of the anonymity of a user within a set. Initial approaches for quantifying user anonymity assumed protocols were deterministic in nature, and used the concept of an anonymity set [2,7,17]. Since then, there has been a large body of research into the construction of metrics for quantifying anonymity. In particular, work has focused on information theoretic metrics capable of analysing probabilistic systems and the relationship between covert channel capacity and anonymity [11,20,21,22,26,33], with a smaller body of work investigating metrics based on combinatorics [13,15].

A definition of the coverage of a metric is important. Clauß and Schiffner generalised all entropy metrics (Shannon entropy, min-entropy, max-entropy) using Reyní entropy, and established a division between global

 $^{^{2}}$ We provide the few facts needed in Appendix A.

metrics that assess the anonymity loss of a system and local metrics that assess the loss of anonymity for an individual user [10]. Toth et al. analyse two entropy-based metrics and their applications in assessing local anonymity [29].

4.1 Using Mutual Information in the Context of Leakage Detection

Assuming the case of discrete random variables, the notation we will use in this paper is as follows. For random variables X and Y, the marginal probability mass function for X is defined as $p_x = \Pr\{X = x\}$. The joint and conditional probability mass functions are similarly defined as $p_{x,y} = \Pr\{X = x, Y = y\}$ and $p_{x|y} = \Pr\{X = x|Y = y\}$ respectively.

We model the side-channel aspect of a web application as an information-theoretic channel; the set of secret user inputs \mathcal{A} are considered as the channel input alphabet, and the observed features \mathcal{O} as the channel output alphabet. The channel conditional probability distribution $D = \Pr\{O = o | A = a\}$ gives us the probability of a side-channel adversary seeing trace feature outputs for a given input choice. The mutual information I(A; O) = I(A; O) = H(A) - H(A|O) gives the *average* reduction in uncertainty about the application inputs A after the adversary has knowledge of the observable outputs O. If A and O are independent, then the mutual information will be zero. When A is completely determined by O, the mutual information will be at its maximum value of H(A).

Input distribution. To compute the mutual information we require a prior distribution for the secret inputs $Pr\{A = a\}$. To accurately estimate information leakage in a real-life scenario, a web application owner may wish to use an input distribution itself estimated from the usage habits of (possibly a subset of) the user population. For example, the medical conditions common for South American users of an international health application are likely to prove to be different for those from East Asian users; two leakage results using distributions calculated from both populations will be more informative and useful than one derived from a single, unrepresentative generic distribution.

Capacity. Alternatively, a worst-case estimate for the information leakage can be computed using the *capacity* of the channel. In essence, the computation of the channel capacity finds the input distribution that leaks the most information across into the trace output distribution. The capacity is defined as follows:

$$C(D) = \max_{\Pr\{A\}} I(A; O).$$

Leakage Detection vs. Exploitation The relationship in general between side-channel information and security is a very complex one. Standaert et al. study the relationship between conditional entropy and the success rates of an adversary in the context of side-channel key-recovery attacks [27], and come to the conclusion that in general, the relationship between the success of an adversary and the conditional entropy of a leakage distribution cannot be fully determined. In our web application context we face a similar problem; the exact structure of the leakage distribution is unknown to the designer as well as the adversary, and, without access to a computationally unbounded sampling phase, cannot be perfectly approximated. As a result we cannot make any definitive conclusions about the success rates of a Bayesian adversary directly from the mutual information or capacity of our application alone. In practice, to be able to concretely assess the true vulnerability of the application it becomes necessary that a leakage evaluation should always be complimented by a separate security evaluation of appropriately defined adversaries.

Nevertheless, mutual information is still an appropriate tool for leakage detection: recent progress ([6]) has been made in constructing a sound test for distinguishing between zero and non-zero mutual information. This distinction between zero and non-zero mutual information serves as a highly useful decision criterion in the engineering problems of identification and mitigation of leaks. In particular, we place significance on this distinction between zero and non-zero mutual information in our design for a leak detection test in section 5. Finally we note that in our web application context, we are typically less concerned with information leaks for a particular instance of a secret input, and the input space is often very small. It therefore makes sense to evaluate web applications using a *global* security metric. In this context, a local metric such as min-entropy gives us the worst-case information leak for a particular input value, but does not give us any further information about the information revealed in the rest of the system. The Shannon entropy computes the average loss of information in the system, and hence is likely to be more suitable in capturing the application performance as a whole.

4.2 Leak Detection Testing Procedure

In practice, the leakage distribution D is unknown to the application designer, and has to be estimated by statistical sampling. When the web application has been sampled n times we obtain an estimated distribution $\hat{D}_n = \hat{\Pr}\{O = o | A = a\}$. In turn, this means we are estimating the mutual information $\hat{I}_n(A; O)$, and similarly will produce an estimate of the capacity $\hat{C}_n(\hat{D}_n)$. The variables we require or produce as part of a side-channel leakage quantification test are detailed in Table 1.

Table 1. Variables required and values computed during a test

A	the web application input(s) we wish to assess for information leakage
0	the set of trace features produced as a result of executing the web application with inputs ${\cal A}$
n	the number of samples taken from the web application
$\hat{D}_n = \hat{\Pr}\{O = o A = a\}$	the conditional probability distribution estimated from n samples
$\hat{I}_n(A;O)$	the estimated mutual information for the application inputs and features, using n samples
$\hat{C}_n(\hat{D}_n)$	the estimated capacity computed using the estimated conditional probability distribution

The quality of the estimation of the probability distribution D is now important, as it introduces a source of error into the estimation of the mutual information and capacity values. Several techniques exist for distribution estimation. The simplest non-parametric approach is the histogram technique; more complicated non-parametric techniques are kernel density estimation, vector quantization, and strategies based around clustering of the data. Parametric approaches under the presence of assumptions could improve the quality of the estimation, but are unlikely to be viable in this context. We note that the importance of the estimation quality here has two-fold implications. Assuming the selection of a suitable estimator, if a designer is unable to sample enough data to approximate the distribution sufficiently well in a particular amount of computational time, then any adversary bounded by the same level of computational power cannot either. This indicates the support size for a successful attack on a web application is a useful consideration.

In [6] Chatzikokolakis et al. find a confidence interval for the estimation of mutual information and capacity over a channel as outlined here. Let I = |A| and J = |O| be the sizes of the input and output domains. A result of their work is that when the true mutual information I(A; O) = 0, for large *n* the value $2n\hat{I}(A; O)$ is approximately χ^2 distributed with (I - 1)(J - 1) degrees of freedom, with a mean of (I - 1)(J - 1)/2n and variance $(I - 1)(J - 1)/2n^2$.

Alternatively, when I(A; O) > 0 the estimate is instead approximately normally distributed, with a mean of $I(A; O) + (I-1)(J-1)/2n + O(1/n^2)$ and a variance v (see Theorem 6.1 in [6]). The approximate bias of the estimator is (I-1)(J-1)/2n, and the authors show that the variance of the estimator in both the zero and non-zero cases is bounded above by IJ/n.

This result means that the accuracy of the estimates can be increased by taking a number of samples larger than the product of the sizes of the input and output domains; keeping the estimate variance low. The ability of the designer to sample enough data to reduce the variance of the test to a reasonable level therefore indicates the applicability of this test in a particular setting. Our web application context is characterised by low-entropy input spaces, and common trace features will typically have fairly small ranges, so it is likely that in most cases the designer will be able to do this. Defining the estimator result as $e = \hat{I}(A; O)$, let $pe = \max(0, e - (I-1)(J-1)/2n)$ be the point estimate for the mutual information. Following Chatzikokolakis et al., if 2ne is inside the 95% confidence interval of the $\chi^2((I-1)(J-1))$ distribution then the confidence interval for the capacity or mutual information is:

$$(0, pe + 1.65\sqrt{v})$$

and if 2ne is *outside* the 95% confidence interval of the $\chi^2((I-1)(J-1))$ distribution then the confidence interval is:

$$(pe - 1.96\sqrt{v}, pe + 1.96\sqrt{v})$$

where v is the variance of the estimated mutual information as defined previously.

A key point to make is that these particular confidence intervals are defined only in the case of *discrete* trace features. In the context of web applications, packet arrival times are likely to be the most common continuous feature containing useful side-channel information. For such continuous features, there are two options:

- Use bucketing strategies to treat the continuous data as discrete. Due to the data processing inequality³ under a bucketing strategy the mutual information estimate will be a lower bound on the true value.
- Compute an estimate of the continuous mutual information⁴ (see Appendix A for a definition).

A cautionary note in the continuous case is that finding good estimators for mutual information is difficult; the optimal estimator is dependent on the (unknown) underlying structure of the data, and there is no universal rate of convergence of an estimator to the true value [24]. Chothia and Guha have devised a statistical hypothesis test for mutual information in the case of discrete inputs and continuous outputs by establishing a rate of convergence for the estimator in the case of zero mutual information [9]. Their result guarantees a good convergence rate in the case of two possible inputs only, but their test is demonstrated to show good performance in the case of multiple inputs.

Multivariate extensions. If more than one feature is to be included in a leak quantification, then there are two options. Denoting the i^{th} feature as O_i :

- 1. Compute the mutual information between the inputs A and n trace features considered jointly: $I(A; (O_1, \ldots, O_n));$
- 2. Compute the multivariate mutual information between the inputs A and n trace features: $I(A; O_1; O_2; \ldots; O_n)$.

The definition of multivariate mutual information I(X;Y;Z) between the triple ensemble of three random variables X, Y and Z essentially describes the gain or loss of information from the set of two variables given knowledge of the third. The definition can be extended to any number of variables and is defined inductively using the definition of conditional mutual information. The conditional mutual information I(X;Y|Z) considers the expected value of I(X,Y) given the value of Z. In terms of probability mass functions, it is defined as:

$$I(X;Y|Z) = \mathbb{E}_{Z}(I(X;Y|Z)) = \sum_{z \in Z} p_{z} \sum_{y \in Y} \sum_{x \in X} p_{x,y|z} \log \frac{p_{x,y|z}}{p_{x|z} p_{y|z}}$$

Multivariate mutual information is defined inductively using this conditional mutual information. For n + 1 random variables X_1, \ldots, X_{n+1} :

$$I(X_1; \ldots; X_n; X_{n+1}) = I(X_1; \ldots; X_n) - I(X_1; \ldots; X_n | X_{n+1}).$$

Figure 5 shows the relationship of these two approaches. Whilst both definitions of mutual information are equally viable, the interpretation of 'true' multivariate mutual information is perhaps less intuitive, is more computationally expensive to compute, and there is no result to our knowledge on bounding the estimate for multivariate mutual information. The alternative definition using the features considered jointly captures more information about the inputs, so with these properties in mind we suggest computing the mutual information between the application inputs and trace features considered jointly.

 $[\]overline{I(f(X);g(Y))} \leq I(X;Y)$ for any random variables X and Y and any functions f and g on the range of X and Y.

⁴ Formally, we require a mutual information definition for one discrete and one continuous random variable.



Fig. 5. An information-theoretic representation of the two approaches for estimating leakage resulting from the consideration of two trace features.

5 The Leak Detection Test illustrated with an Example

Using our leakage metric proposal, in the following section we describe a leak detection test capable of identifying the *presence* of many common information leaks.

Let $\mathbf{P} = \{p_1, \ldots, p_n\}$ be a trace consisting of n packets sampled from a web application. Each trace is reduced to a corresponding vector of features $\mathbf{O} = \{\mathbf{o_1}, \ldots, \mathbf{o_n}\}$ with each $\mathbf{o_i}$ being a feature vector of length k representing the values of a set of k selected features considered jointly for the packet p_i at position i in the trace. Defining N as the length of the longest trace, we construct a set O^i that contains the feature values at the i^{th} position in each of the sampled traces, for $i \in \{1, \ldots, N\}$. For example, O^3 might contain the set of sizes of the third packet in each sampled trace. We then estimate the conditional probability distribution $\hat{D}^i(O^i, A) = \hat{\Pr}\{O^i = o | A = a\}$ of the selected packet features O^i at each possible position in a trace, for $i \in \{1, \ldots, N\}$. Finally, using the estimated distributions we estimate mutual information values $\hat{v}^i = \hat{I}(A; O^i)$ between the application inputs A and the packet features O^i . Using the test of Chatzikokolakis et al., we then compute corresponding 95% confidence intervals for zero and non-zero mutual information for each of the $i \in \{1, \ldots, N\}$ estimates.

We identify leaks by checking the confidence intervals for zero mutual information: an estimated mutual information value at position i that falls outside of the 95% confidence interval for zero mutual information can be taken as evidence of an information leak at position i. A positive result at a particular position will not necessarily describe the true size of the information leak; a confidence interval for the point estimate of non-zero mutual information that is further away from the corresponding interval for zero mutual information suggests a larger information leak than a confidence interval for the point estimate that is closer to the confidence interval for zero mutual information, but does not guarantee it. Careful interpretation of the leak detection results is paramount, and a security analysis to complement an information-theoretic analysis is always recommended.

A common example of a detectable information leak in practice occurs when two subsets of the application inputs each redirect the user to a differently sized web page. The leak detection test will identify the information leak towards the *end* of the subsection of the trace that corresponds to the server response; given the sequence of packets that correspond to the server returning the new web page, the initial packets are all likely to be equivalent in size, typically of the maximum segment size (MSS) for the TCP session, but the length of the sequence is likely to be longer, or the final packet larger, for the larger web page.

The left-hand side of Figure 6 shows a graph of mutual information estimates and corresponding zero and non-zero confidence intervals extracted from a particular state within our online health application, using the presence of a TCP acknowledgement flag as a packet feature. The state consists of a 'Yes or No' style question, with the user being redirected to different web pages depending on their choice. One web page is slightly larger in size due to the inclusion of additional text on the page. The highest peak in the graph occurs at position 22

in the sequence. The choice that leads to the larger web page results in extra packets being downloaded, and as a result a larger amount of acknowledgement packets are present in traces for that choice. A conclusion to draw here is that in this circumstance, an information leak is present that does not take into account packet sizes at all, illustrating that the depth of potential side-channel attacks on web applications extends beyond an analysis of packet sizes.



Fig. 6. Mutual information estimates at each index in a sequence of packets, (left) using the presence of a TCP acknowledgement flag as a packet feature, and (right) using the packet size as a packet feature. There are two possible inputs to the application at this state, with the maximum entropy for these two inputs represented by the horizontal line.

Multivariate extensions. It is possible that a detection test considering additional packet positions jointly (estimating $I(A; O^i, \ldots, O^j)$) may detect an information leak that a univariate test could not: $I(A; O^i) = 0$ and $I(A; O^j) = 0$, but $I(A; O^i, \ldots, O^j) > 0$. The identification and selection of the additional positions is a non-trivial task that will typically require additional domain knowledge from the designer. An exhaustive multivariate test is computationally feasible when considering the joint distribution of a small number of packet positions; for instance, given a trace of length n, we can check all possible pairs of packets in $O(n^2)$ time.

A second consideration is whether a combination of packet features (e.g. combining packet size and direction) could result in the identification of an information leak that would not be detected by an individual analysis of packet size and direction. Again, additional knowledge from the designer is likely to be required in identification of appropriate feature combinations, but the cardinality of the set of possible packet features is not particularly large; perhaps in the order of 10 or so, and the combination of some of these features will not always make sense, so it may be possible to exhaustively search the joint feature space. A general benefit of mutual information as our proposed metric choice is that it extends quite naturally into the multivariate setting and is applicable in these sorts of scenarios, whereas other metrics are not necessarily quite as easily extended.

6 Using the Test in the Presence of Countermeasures

Countermeasures against side-channel attacks on web applications vary in terms of how well they are supported by servers or client software, have different efficiency overheads and the effectiveness of many is dependent on the powers of a considered adversary. In general a rough rule-of-thumb is that an adversary with visibility of communication at a particular network layer will be able to subvert countermeasures applied at that layer. Many of the countermeasures have been suggested in the context of defending against website fingerprinting attacks, and are suitably generic in their approach as a result. In our web application context, it is likely that with a thorough leakage analysis, countermeasures can be applied in a more targeted, application-specific fashion, and hence potentially at a smaller cost in overhead. In Table 2 we provide a tabular representation of some of the most popular transport and application-layer countermeasures as considered in the research community and the packet/trace features affected most by the countermeasure. A more detailed discussion of the effectiveness of the countermeasures is beyond the scope of this paper, but can be found in the referenced publications.

Countermeasure	OSI Layer	Main features affected
HTTP header padding[19,28]	Application	Packet sizes, no. of packets
HTTP Range property[19,28]	Application	Packet sizes, no. of packets
HTTP Pipelining[19,28]	Application	No. of packets, web object sizes
Injecting useless requests[19]	Application	No. of packets, trace size
Background loading of random page[23,28]	Application	No. of packets, packet sizes & timings
Object prefetching[28,31]	Application	No. of packets, trace size
TLS/IPSec padding[12]	Presentation/Network	A Packet sizes
Direct traffic sampling[30]	Transport	No. of packets, packet sizes & timings
Traffic morphing[28,30]	Transport	No. of packets, packet sizes & timings
Extending TCP Option field[19,12]	Transport	Packet sizes
TCP Maximum Segment Size (MSS) renegotiation[19]	Transport	Packet sizes, no. of packets
TCP advertising window[19]	Transport	Packet sizes, no. of packets
Forced TCP retransmissions[19]	Transport	No. of packets, web object size
Delaying requests and ACK responses[19]	Transport	Packet arrival times

 Table 2. Proposed transport and application-layer countermeasures for mitigation of side-channel attacks on web applications and website fingerprinting attacks

In the following section we demonstrate the usefulness of applying our quantification metric and leak detection strategy in an analysis of countermeasure effectiveness, using our online health application example. We use our metric to detect and quantify the information leaked through packet sizes in the presence of frequently proposed countermeasures. Our aim is not to provide a comprehensive analysis of every plausible countermeasure; rather we demonstrate the applicability of our strategy in general.

Several countermeasures require modifying traffic emanating from the host web server, and so we are unable to directly apply those countermeasures when testing live web applications we do not control. Additionally, the applicability of several countermeasures is dependent on the particular web server software hosting the application. As a result, and as we are not attempting an exhaustive comparison of all countermeasures, we constructed two distinct countermeasures that mitigate information leaks within packet sizes, and *simulated* their application on real sampled data. The following list outlines our choices and our approach for their simulation:

- Individual packet padding: for each captured packet, we artificially increase its size. The size can be increased by a random amount, or can be fixed to a pre-determined amount. In practice, the random strategy can be achieved by padding the TCP options field uniformly at random, by re-negotiating the TCP maximum segment size (MSS) for a session, or by randomly increasing the size of requests and responses at the application layer. The fixed-sized strategy is more difficult to apply in practice, but is an interesting option for comparison. We simulate both approaches on our sampled packets: in the random approach, we retrieve the TCP header and randomly add bytes to the option field, and in the fixed-size approach we artificially increase the packet size to the Ethernet Maximum Transmission Unit (MTU) size. Intuitively, this countermeasure can be expected to 'shrink' the vector of mutual information estimates, but is limited by being restricted to modifying packets that are present only.
- **Insertion of fixed-size dummy packets**: for each sampled trace, we insert a random amount of fixedsize dummy packets into the server response at various positions in the sequence. An obvious strategy at the application layer for obfuscating information leaks resulting from varying server response sizes is to add random noise to the response. In practice this can be achieved by re-downloading data at random, pipelining, injecting useless requests, padding HTTP headers, or even by the application itself generating

randomly-sized data in response traffic. The expected effect on our traces should again be to shrink the vector of estimates, by in some sense shuffling the order of the packets within a trace.

To aid in an analysis of their efficiency, the effects of these two countermeasures are deliberately intended not to overlap. The first countermeasure affects traces in the vertical direction, by distorting the properties of individual packets, and the second acts in the horizontal direction by diffusing the distributions at each packet position across into neighbouring positions. In reality, the most effective side-channel countermeasure is likely to be one that affects traces in both these ways. Exempting the case of object pre-fetching, all countermeasures designed to effect leaks resulting from the size of an object within a trace can clearly only *increase* the size of a trace. Since the majority of such leaks are likely to result from systematic differences in the total size of the server response, countermeasures have the choice of making the distributions of the responses indistinguishable by attempting to make the response sizes the same, or by attempting to make the response sizes sufficiently random. The application of our simulated dummy packet countermeasure follows the latter paradigm; by inserting a random number of packets into the trace in the appropriate locations, the countermeasure attempts to make the size distributions under each input less distinguishable.

6.1 Results with regards to Countermeasures

We tested the same online health application state we used to describe our leak detection test previously, and consider a scenario in which a user connects to the application using a normal Internet connection protected only by HTTPS, with the adversary able to capture packets at the transport-layer (the adversary can 'see' TCP and TLS headers). To test the application, we sampled the state repeatedly using automated browser testing software called Selenium [25], and captured packets using a native pcap library. We first ran our leak detection test without any countermeasures applied, using packet sizes as a feature. The results are shown in the right-hand side of Figure 6. An analysis indicates that there are two leak sources present that result from information contained in packet sizes: the size of the first packet, sent by the client, is extremely distinguishing, and the server response towards the middle of the trace is also strongly distinguishing.

Figure 7 shows the results of our leak detection test re-run in the presence of our four countermeasures applied individually, and Table 3 gives the mean increase in size of the sampled traces as a result of the application of each countermeasure.

Table 3. The average size increase of captured traces in the presence of the four considered countermeasures.

Countermeasure	Mean size overhead (%)
Randomly padding TCP option field	1.2
Padding each packet to the Ethernet II MTU size	95.4
Randomly inserting 10-15 dummy packets into all traces	48.1
Joint combination of insertion and random padding strategies	49.9

We simulated a limited packet-padding countermeasure by emulating the TCP option field padding strategy; the TCP option field can contain at most 40 bytes, so we pad each packet uniformly at random in the range of 0 to 40 bytes. We can see that this countermeasure has a very strong effect on the leak resulting from the size of the first packet in the trace, but a negligible effect on the leak resulting from the size of the server response. An analysis of the source code of the application reveals that the difference in size between the two possible client requests is slightly smaller than the maximum size of the TCP option field; hence the small amount of noise added is sufficient to disguise the information leak in this case. The difference in size of the packets between positions 20 and 30 is typically larger than 40 bytes, and hence the countermeasure is insufficient to disguise the leak from the server response. Finally, the countermeasure cannot disguise the fact that traces are slightly longer when the user selects the input that corresponds to the larger web page, and we can see that peaks in the graph at the final few packet positions are indicative of this information leak.



Fig. 7. Leak detection results analysing the distribution of the sequence of packet sizes, in the presence of (top-left) randomly padding the TCP option field, (top-right) padding each packet to the Ethernet MTU, (bottom-left) adding randomly-sized dummy packets to the server response and (bottom-right) combining dummy packet insertion and random individual packet padding strategies

To simulate the fixed-sized packet-padding approach, we artificially set the size of each packet to the Ethernet II MTU of 1500 bytes. This countermeasure has a much larger overhead than the padding of TCP option fields (95.4% to 1.2%), but eliminates the information leak contained in the first packet and from the server response packets in the middle of the sequence. Again, the individual packet padding strategy employed here suffers from being unable to disguise the longer traces for one of the inputs, and so the end few positions of the packet sequence show a significant information leak.

Modifying the server response by adding between 10 and 15 dummy packets into the trace has a noticeable reductive effect on the size of the leak. An increased amount of the point estimates are now within the 95% confidence intervals for zero information leakage. However, in the middle and towards the end of the trace we can see that there are still packet positions with estimates noticeably outside of the interval for zero information leakage. A conclusion to draw in this scenario is that the number of packets randomly inserted is both insufficient for nullifying the leak in the middle of the trace, and for nullifying the effect of the larger server response producing longer traces.

Finally, we tested the combination of the TCP option field individual packet padding strategy and the insertion of dummy packets jointly. We first inserted the dummy packets, and then applied the random padding. The reductive effect on the mutual information estimates corresponding to the server response is clearly the best so far. To determine whether this level of protection is sufficient, the designer should consider their security goals with these results and with the results from a separate security analysis. Whilst the translation of our particular simulated countermeasures into a practical implementation is not straightforward, these results are indicative of how our detection test can be used to analyse the effectiveness of noise-based side-channel countermeasures.

7 Conclusion

We have discussed the challenges involved in the detection (and quantification) of information leaks in the context of side-channel attacks on web applications, and focused on dissimilarities between evaluating the strength of an adversary and quantifying the strength of an information leak. In our approach we consider a web application as an information theoretic channel, with the application inputs considered as the channel inputs, and features extracted from network traces as the channel outputs. Under this model, we have proposed using an estimate of the mutual information or capacity of the channel as a metric for detecting instances of information leakage in a web application. Our approach has the benefit of a degree of statistical rigour, can cope with non-deterministic traces, and does not require application-specific distance metrics or thresholds for classifications.

We use our proposal to construct a leak detection test for the identification of information leaks, and present an example application of it in analysing potential side-channel analysis countermeasures using a popular online health website. The leak detection test outlined here will aid in the automatic detection of the majority of information leaks from web applications. To accurately quantify the size and implications of the information leak detected by the test, additional domain knowledge from the application designer is likely to be required. An interesting research challenge is to construct methods that detect leaks by mechanically and efficiently analysing combinations of features and packet positions over the entire feature space. This problem is very similar to the challenge of feature extraction considered in the field of machine learning, and so further research in this direction may yield useful results.

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A Basic information theory

Assuming the case of discrete random variables, the notation we will use in this paper is as follows. For random variables X and Y, the marginal probability mass function for X is defined as $p_x = \Pr\{X = x\}$. The joint and conditional probability mass functions are similarly defined as $p_{x,y} = \Pr\{X = x, Y = y\}$ and $p_{x|y} = \Pr\{X = x|Y = y\}$ respectively.

Shannon entropy. Let X and Y be two discrete random variables, supported on \mathcal{X} and \mathcal{Y} respectively. The Shannon entropy of X is defined as the *average* uncertainty about X. Mathematically it is defined as:

$$H(X) = -\sum_{x \in \mathcal{X}} \Pr\{X = x\} \log \Pr\{X = x\},\$$

and when the logarithm is taken with base 2 measures the remaining uncertainty about X in bits. The conditional entropy H(X|Y) gives the remaining average uncertainty about the variable X when the value Y has taken is known, and is defined similarly as

$$H(X|Y) = -\sum_{y \in \mathcal{Y}} \Pr\{Y = y\} \sum_{x \in \mathcal{X}} \Pr\{X = x | Y = y\} \log \Pr\{X = x | Y = y\}.$$

Mutual information. The *mutual information* I(X;Y) between X and Y computes the information gained about X if we observe Y, and is related to the marginal and conditional entropies as follows:

$$I(X;Y) = H(X) - H(X|Y).$$

If X and Y are independent, then the mutual information will be zero. When X is completely determined by Y, the mutual information will be at its maximum value of H(X).

Channels. An information-theoretic channel takes inputs from an input alphabet \mathcal{X} and sends across the channel outputs from a set of channel outputs \mathcal{Y} . The channel itself is described by the probability distribution Q, where $Q(y|x) = \Pr\{Y = y|X = x\}$ gives the probability of seeing a particular output given a particular input. The mutual information I(X;Y) describes the amount of information that can be sent over the channel. Contextually this gives us the average amount of side-channel information available to an adversary.

Note that the input probability distribution affects the mutual information result. Hence, the worst-case information leakage occurs when the amount of information sent over the channel is maximised. The channel *capacity* is defined as the mutual information of the channel maximised over all possible input distributions:

$$\operatorname{Capacity}(Q) = \max_{\Pr\{X\}} I(X;Y).$$

The capacity can be approximated to within a specified error using the iterative Blahut-Arimoto algorithm [1,4].

Other definitions of entropy. The Shannon entropy as described above gives the *average* uncertainty about a random variable; however alternative definitions exist. The min-entropy, defined as $H_{\infty}(X) = -\log \max_{x \in X} \Pr\{X = x\}$, gives the amount of uncertainty needed to be overcome to be able to identify the most likely element only. Shannon entropy and min-entropy are both specific instances of Reyní entropy:

$$H_{\alpha}(X) = \frac{1}{1-\alpha} \log \left(\sum_{x \in X} \Pr\{X = x\}^{\alpha} \right).$$

Intuitively, we can see that if $\alpha = 1$, then Reyní entropy is equivalent to Shannon entropy, and as α approaches infinity, Reyní entropy tends to min-entropy. One can also see that increasing the parameter α can aid in eliminating outliers from the entropy calculation. Clauß and Schiffner discuss the applicability of these metrics in anonymity-scenarios in [10].

Continuous mutual information. In the context of discrete inputs and continuous outputs we require a slightly different construction for mutual information. For a discrete random variable X and a continuous random variable Y, the equation for mutual information becomes

$$I(X;Y) = \sum_{x \in X} \int_{Y} p(x,y) \log\left(\frac{p(x,y)}{p(x)p(y)}\right) dy.$$