

Code-based signatures without trapdoors through restricted vectors

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Abstract. The Schnorr-Lyubashevsky approach has been shown able to produce secure and efficient signature schemes without trapdoors in the lattice-based setting, exploiting Gaussian distributions in the Euclidean metric and rejection sampling to tune the signature probability distributions. Translating such an approach to the code-based setting has revealed to be challenging, especially for codes in the Hamming metric. In this paper, we propose a new adaptation of the Schnorr-Lyubashevsky framework to codes in the Hamming metric exploiting restricted vectors, which allows avoiding existing attacks. We provide some preliminary arguments to assess the security level of the new scheme and for computing the relevant parameters. We show that the new scheme achieves compact keys and signatures even without considering structured codes.

Keywords: Code-based cryptography · digital signatures · post-quantum cryptography

1 Introduction

There are basically two approaches to code-based digital signatures. The first one is derived from the “hash-and-sign” paradigm used for instance to achieve RSA signatures, and encounters some difficulty when applied to the code-based setting. This is due to the difficulty of randomly generating a decodable syndrome, yielding code-based schemes that are inefficient or insecure (or both). Two historical proposals in this line are CFS [13] and KKS [21], which however have severe limitations. In fact, it is very difficult to find secure though efficient instances of KKS [24]. Also CFS requires some extreme choices of the code parameters to be efficient (e.g., very high code rates), and this exposes the system to Goppa code distinguishers [18]. Some variants of CFS aimed at using non-algebraic codes and reducing the public key size have been proposed [7], but changing the underlying family of codes yielded to successful cryptanalysis [27].

Another important drawback of existing code-based signature schemes relying on the hash-and-sign paradigm is the large size of the public keys. A recent and relevant scheme in this line, Wave [14], introduces a new approach relying on the hardness of decoding vectors with very large weight, and achieves a

public key size growing quadratically with the security level, which is an important improvement over CFS. Nevertheless, the public key size in Wave is over 3 megabytes for 128-bit security, which is still rather large.

A different approach to code-based signatures, which has the advantage of not relying on any trapdoor for key derivation, is that of applying the Fiat-Shamir transform [19] to a code-based identification scheme. In fact, consolidated zero-knowledge code-based identification schemes exist since a long time [31], which however exhibit significant soundness errors and thus require many repetitions. This leads to large signature sizes when these schemes are used to achieve digital signatures. Subsequent variants of these schemes aim at overcoming such limitations [33, 12, 1, 17, 9, 10, 5], but their characteristics are still far from being comparable with those of signature schemes relying on other mathematical objects than codes, like lattices. One of the main advantages of lattice-based schemes is that they can exploit the approach introduced by Lyubashevsky in [23], achieving very compact keys and short signatures, besides high algorithmic efficiency. Such an approach is at the basis of Dilithium [16], one of the most promising digital signature schemes participating to the ongoing NIST competition.

This motivates many attempts to translate the Schnorr-Lyubashevsky approach into the domain of code-based schemes, as done in [25, 30, 22]. In most cases, however, these attempts resulted in a successful cryptanalysis of the corresponding schemes [29, 15, 2, 6]. While there are adaptations of the Schnorr-Lyubashevsky approach in the rank metric code-based setting that are still considered safe [3] and achieve reasonable performance, no valid solution has been found in the Hamming metric code-based setting to date. In many of the aforementioned examples using codes in the Hamming metric, binary codes and sparse signatures were used, obtained from a sparse secret key via linear algebra: this feature is at the core of the corresponding attacks and definitely represents a weak choice. The crucial difference between codes and lattices lies in how a small vector can be defined. For the Hamming metric, smallness has normally been associated with sparsity, thus requiring the existence of a large number of zero entries in the noisy vectors: this unfortunately makes the noise ineffective in masking the secret key. Lattices instead are defined in the Euclidean metric, for which a small vector does not necessarily contain zero entries: this is the key to dispose of small though secure noisy vectors.

We propose a novel code-based adaptation of the Lyubashevsky signature scheme in the Hamming metric. Inspired by [14] and [5], we avoid sparsity by using codes defined over a large finite field and secret keys formed by dense vectors whose entries take values in a restricted subset of the possible values, which we call *restricted vectors*. By doing this, we base the one-wayness of the key generation algorithm on the recently introduced Restricted Syndrome Decoding Problem (R-SDP), which has been proven to be NP-complete [5]. Similarly to the Schnorr-Lyubashevsky approach, we employ a high noise term to properly hide the secret key, and exploit rejection sampling to tune the obtained distribution and avoid signatures that may leak information about the secret key.

2 Notation

For two integers a and b , we denote as $[a; b]$ the set of integers x such that $a \leq x \leq b$; if a and b are instead reals, we denote as $\llbracket a; b \rrbracket$ the range defined by all reals x such that $a \leq x \leq b$. We denote as \mathbb{F}_q the finite field with q elements. We consider the case of q prime, and sometimes represent the elements of the field as $[-\lfloor q/2 \rfloor; \lfloor q/2 \rfloor]$. We denote matrices and vectors with bold capital and small letters, respectively. Given a matrix \mathbf{A} , we denote its entry in the i -th row and j -th column as $a_{i,j}$; for a vector \mathbf{a} , its i -th entry is indicated as a_i . The identity matrix of size r will be indicated as \mathbf{I}_r . By support of a vector \mathbf{a} , we mean the set containing the indexes of non-zero coordinates. For two vectors \mathbf{a} and \mathbf{b} having the same length n , we indicate their inner product as

$$\langle \mathbf{a}; \mathbf{b} \rangle = \sum_{i=0}^{n-1} a_i b_i.$$

If \mathcal{D} is a probability distribution, we write $a \sim \mathcal{D}$ if a is a random variable distributed according to \mathcal{D} . For $\mathcal{D} : \mathbb{F}_q \mapsto \llbracket 0; 1 \rrbracket$, with some abuse of notation, the expression $\mathbf{a} \sim \mathcal{D}$ means that each entry of $\mathbf{a} \in \mathbb{F}_q^n$ is distributed according to \mathcal{D} . If \mathcal{D} is a discrete probability distribution, we denote as $\mathcal{D}(\mathbf{a})$ the probability that \mathcal{D} outputs \mathbf{a} . For a set A , we write $a \stackrel{\$}{\leftarrow} A$ if a is uniformly picked at random among all the elements of A .

2.1 Coding theory preliminaries

Let \mathcal{C} denote a linear code over \mathbb{F}_q with length n , dimension k , redundancy $r = n - k$ and rate $R = k/n$. We represent codes through their parity-check matrix, i.e., a full rank matrix $\mathbf{H} \in \mathbb{F}_q^{r \times n}$ such that $\{\mathbf{H}\mathbf{c}^\top = \mathbf{0} \mid \forall \mathbf{c} \in \mathcal{C}\}$. A *systematic* parity-check matrix is a parity-check matrix in the form $\mathbf{H} = [\mathbf{I}_r, \mathbf{P}]$, where $\mathbf{P} \in \mathbb{F}_q^{r \times k}$. For $\gamma \leq \lfloor q/2 \rfloor$, we denote with $S_{\gamma,t}$ the set of length- n vectors with entries over $\{0, \pm 1, \dots, \pm \gamma\} \subseteq \mathbb{F}_q$ and support size t . For the set of vectors with support size not greater than t , we instead write $B_{\gamma,t} = \bigcup_{i=0}^t S_{\gamma,i}$. The Restricted-Syndrome Decoding Problem (R-SDP) is defined as follows [5].

Problem 1. R-SDP $_{\gamma, \leq t}$: R-SDP with bounded support size t

Let $\mathbf{H} \in \mathbb{F}_q^{r \times n}$, $\mathbf{s} \in \mathbb{F}_q^r$ and $t \in \mathbb{N}$. Find $\mathbf{e} \in B_{\gamma,t}$ such that $\mathbf{H}\mathbf{e}^\top = \mathbf{s}$.

As proven in [5] with a reduction from the decoding problem in the Hamming metric, the decisional version of the above problem is NP-complete, regardless of the value of γ . In this paper we also consider a slightly modified version of Problem 1, where we require the support of the searched vector to be exactly t (and hence, require the solution vector to be in $S_{\gamma,t}$). We denote the associated problem as R-SDP $_{\gamma,=t}$. It is easily seen that also the decisional version of R-SDP $_{\gamma,=t}$ is NP-complete. In fact, any polynomial time solver for R-SDP $_{\gamma,=t}$ can be used to solve R-SDP $_{\gamma, \leq x}$ for a given x , by invoking it for no more than x times. This trivially shows that such an efficient solver cannot exist. Notice that the difference between these two problems is mostly formal, since they can be solved with the same techniques. More details are given in Section 5.

3 New signature scheme

The scheme we propose is parameterized by the positive integers $r, n, q, b, w_E, w_c, t_E$, with $r, b, t_E < n$ and $w_E, w_c \leq b$. It additionally employs two probability distributions $\mathcal{D}_y, \mathcal{D}_z$ defined over \mathbb{F}_q^n . Finally, we make use of an hash function Hash that outputs vectors of size b with entries over $\{0, \pm 1\} \subseteq \mathbb{F}_q$ and support of size w_c . The scheme we propose consists of the following triplet of algorithms.

Key generation

1. Select at random $\mathbf{P} \in \mathbb{F}_q^{r \times k}$ and set $\mathbf{H} = [\mathbf{I}_r | \mathbf{P}]$.
2. Select at random $\mathbf{E} \in \{0, \pm 1\}^{b \times n}$ such that each column has support size w_E , and each row has support size not lower than t_E ;
3. Compute $\mathbf{S} = \mathbf{E}\mathbf{H}^\top \in \mathbb{F}_q^{b \times r}$;
4. Set $\text{sk} = \mathbf{E}$, $\text{pk} = \{\mathbf{H}, \mathbf{S}\}$.

Signature generation On input a message m :

1. Sample $\mathbf{y} \in \mathbb{F}_q^n$ from \mathcal{D}_y ;
2. Compute $\mathbf{s}_y = \mathbf{y}\mathbf{H}^\top$;
3. Compute $\mathbf{c} = \text{Hash}(m, \mathbf{s}_y)$;
4. Compute $\mathbf{z} = \mathbf{c}\mathbf{E} + \mathbf{y}$;
5. Perform rejection sampling to tune the distribution of \mathbf{z} to \mathcal{D}_z ;
6. Output $\sigma = \{\mathbf{z}, \mathbf{c}\}$.

Signature verification On input m and $\sigma = \{\mathbf{z}, \mathbf{c}\}$

1. Verify that $\mathcal{D}_z(\mathbf{z}) \neq 0$, reject otherwise;
2. compute $\mathbf{s}_y = \mathbf{z}\mathbf{H}^\top - \mathbf{c}\mathbf{S}$;
3. accept if $\mathbf{c} = \text{Hash}(m, \mathbf{s}_y)$, reject otherwise.

The public key size corresponds to the number of bits one needs to represent \mathbf{H} and \mathbf{S} . Notice that, since \mathbf{H} is random and systematic, it is fully represented through the seed used to generate \mathbf{P} . To represent \mathbf{S} we need $br \lceil \log_2(q) \rceil$ bits, while exploiting an efficient representation for arrays in \mathbb{F}_q , we can use $\lceil br \log_2(q) \rceil$ bits. The signature size is given by the size of \mathbf{z} plus that of \mathbf{c} . To represent \mathbf{z} , we need $n \lceil \log_2(\bar{\gamma}) \rceil$ bits which, using an efficient representation, can be reduced to $\lceil n \log_2(\bar{\gamma}) \rceil$. For \mathbf{c} , it is enough to send its support, together with a single bit for each non-null entry, stating whether it is 1 or -1 : this requires $w_c + n \lceil \log_2(n) \rceil$ bits. We choose \mathcal{D}_y as the uniform distribution over $\{0, \pm 1, \dots, \pm \gamma\} \subseteq \mathbb{F}_q$, and set \mathcal{D}_z as the uniform distribution over $\{0, \pm 1, \dots, \pm \bar{\gamma}\} \subseteq \mathbb{F}_q$, with $\bar{\gamma} < \gamma$. With rejection sampling, we tune the distribution of each entry in \mathbf{z} to be uniformly distributed over $\{0, \pm 1, \dots, \pm \bar{\gamma}\} \subseteq \mathbb{F}_q$, and hence make \mathbf{z} indistinguishable from a uniform element of $S_{\bar{\gamma}, n}$. In particular, we reject each entry with a very low probability ϵ , so that the average number of signatures that one computes before the signing algorithm outputs something is given by $(1 - \epsilon)^{-n}$. The reasons behind the rejection sampling criterion, as well

as recommendations on how to choose γ and $\bar{\gamma}$, will be presented in Section 4. In the verification process, to test whether $\mathcal{D}_z(\mathbf{z})$ is null or not, it is enough to verify that each entry of \mathbf{z} is not outside $\{0, \pm 1, \dots, \bar{\gamma}\}$. Finally, it is easily seen that an honest signature always gets accepted, since

$$\mathbf{z}\mathbf{H}^\top - \mathbf{c}\mathbf{S} = (\mathbf{c}\mathbf{E} + \mathbf{y})\mathbf{H}^\top - \mathbf{c}\mathbf{E}\mathbf{H}^\top = \mathbf{y}\mathbf{H}^\top = \mathbf{s}_y.$$

4 Statistics and rejection sampling

Let us study the statistical distribution of the entries of each produced signature. Due to lack of space, all the proofs of this section are reported in Appendix A. We start by deriving the probability distribution of the entries in the product $\mathbf{c}\mathbf{E}$; to do this, we consider the inner product between a random $\mathbf{c} \in \{0, \pm 1\}^b$ and a vector $\mathbf{e} \in \{0, \pm 1\}^b$ modeling a column of the secret \mathbf{E} .

Lemma 1. *Let $w_E, w_c \in \mathbb{N}$ such that w_E is even and $\min\{w_c, w_E\} < q$. Let $\mathbf{e} \in \{0, \pm 1\}^b$ with support size w_E , such that $w_E/2$ coordinates are equal to 1 and $w_E/2$ coordinates are equal to -1 . Let $\mathbf{c} \stackrel{\$}{\leftarrow} \{0, \pm 1\}^b$, with support size w_c . Then, the probability that $\langle \mathbf{c}; \mathbf{e} \rangle$ is equal to $\beta \in \mathbb{F}_q$ is given by*

$$g_{q, w_E, w_c, b}(\beta) = \sum_{\substack{v=\beta_q \\ v \text{ and } \beta_q \text{ have the same parity}}}^{\min\{w_E, w_c\}} 2^{-v} \frac{\binom{v+\beta_q}{2} \binom{w_E}{v} \binom{b-w_E}{w_c-v}}{\binom{b}{w_c}},$$

where $\beta_q = \min\{\beta, q - \beta\}$.

Lemma 2. *Let $w_E, w_c \in \mathbb{N}$ such that w_E is even and $\min\{w_c, w_E\} < q$. and Let $\mathbf{e} \in \{0, \pm 1\}^b$ with support size w_E , $w_E/2$ coordinates equal to 1 and $w_E/2$ coordinates equal to -1 . Let $\mathbf{c} \stackrel{\$}{\leftarrow} \{0, \pm 1\}^b$ with support of size w_c , and $y \stackrel{\$}{\leftarrow} \{-\gamma, \dots, \gamma\}$. Then, $\Pr[\langle \mathbf{c}; \mathbf{e} \rangle + y = \beta]$ is equal to*

$$\tilde{g}_{q, w_E, w_c, b, \gamma}(\beta) = \frac{\sum_{x=-\gamma}^{\gamma} g_{q, w_E, w_c, b}(\beta - x)}{2\gamma + 1}.$$

In the signing algorithm, we employ a rejection sampling criterion to tune the distribution of produced signatures to a desired target. In particular, we want each entry of \mathbf{z} to follow the uniform distribution over $\{0, \pm 1, \dots, \pm\gamma\} \subseteq \mathbb{F}_q$. To estimate the rejection rate, we rely on the following proposition, which in turn relies on the well-known rejection sampling lemma.

Proposition 1 (Rejection sampling). *Let $w_E, w_c, \gamma, \bar{\gamma} \in \mathbb{N}$ such that w_E is even, $\min\{w_c, w_E\} < q$ and $\bar{\gamma} < \gamma < q$. Let $\mathbf{e} \in \{0, \pm 1\}^b$ with support size w_E , $w_E/2$ coordinates equal to 1 and $w_E/2$ coordinates equal to -1 . Let \mathcal{F} be the uniform distribution over $\{0, \pm 1, \dots, \pm\bar{\gamma}\}$, with probability distribution $f : \mathbb{F}_q \mapsto \llbracket 0; 1 \rrbracket$. Let $M = \max_{\beta \in \mathbb{F}_q} \left\{ \frac{f(\beta)}{\tilde{g}_{w_E, w_c, q, b, \gamma}(\beta)} \right\}$ probability distribution function defined by $f : \mathbb{F}_q \mapsto \llbracket 0; 1 \rrbracket$. Let \mathcal{G} be the distribution resulting from the following experiment:*

1. sample $\mathbf{c} \stackrel{\$}{\leftarrow} \{0, \pm 1\}^b$ with support of size w_c ;
2. sample $\mathbf{y} \stackrel{\$}{\leftarrow} \{0, \pm 1, \dots, \pm \gamma\}$;
3. compute $\mathbf{z} = \langle \mathbf{c}; \mathbf{e} \rangle + \mathbf{y}$;
4. output \mathbf{z} with probability $\frac{f(\mathbf{z})}{M \bar{g}_{w_E, w_c, q, b, \gamma}(\mathbf{z})}$.

Then, \mathcal{G} outputs something with probability $1/M$, and the samples obtained from \mathcal{G} are distributed according to \mathcal{F} .

As a trivial application, we extend the above proposition to the case of multiple samples obtained from \mathcal{G} . In other words, we consider a vector \mathbf{z} obtained by repeating the experiment of \mathcal{G} for n times, i.e.:

1. we choose $\mathbf{E} \in \mathbb{F}_q^{b \times n}$ such that each column has support size w_E and contains w_E entries equal to 1, and -1 entries equal to -1 ;
2. we pick $\mathbf{c} \stackrel{\$}{\leftarrow} \{0, \pm 1\}^b$ with support size w_c and $\mathbf{y} \stackrel{\$}{\leftarrow} \{0, \pm 1, \dots, \pm \gamma\}$;
3. we compute $\mathbf{z} = \mathbf{c}\mathbf{E} + \mathbf{y}$;
4. we output \mathbf{z} with probability $\frac{1}{M^n} \prod_{i=0}^{n-1} \frac{f(z_i)}{\bar{g}_{w_E, w_c, q, b, \gamma}(z_i)}$.

It is easily seen that the resulting distribution outputs something with probability M^{-n} , and that each entry of the sample is distributed according to \mathcal{F} .

5 Solving the R-SDP

The authors of [5] have studied the complexity of solving the R-SDP $_{\gamma,=t}$, for the sole case of $\gamma = 1$ and $t = n$, via techniques borrowed from the ternary SDP [11]. In this section we generalize such an analysis, and consider larger values of γ , as well as values of t that are lower than n . For space reasons, the proofs relevant to this section are reported in Appendix B.

We consider the setting in which \mathbf{H} is the parity-check matrix of a random code \mathcal{C} with length n and dimension k , and assume that at least one solution always exists. For the fixed support size problem (i.e., R-SDP $_{\gamma,=t}$), we assume that the target syndrome is picked as $\mathbf{s} \stackrel{\$}{\leftarrow} \{\mathbf{e}\mathbf{H}^\top \mid \mathbf{e} \in S_{\gamma,t}\}$; in such a case, the number of solutions can be estimated as

$$N_{\gamma,=t} = 1 + \frac{|S_{\gamma,t}| - 1}{q^{n-k}} \approx 1 + \binom{n}{t} 2^{t(1+\log_2(\gamma)) - (n-k)\log_2(q)}. \quad (1)$$

Indeed, consider that for each $\mathbf{e} \in S_{\gamma,t}$, we have that $\mathbf{e}\mathbf{H}^\top$ is random over \mathbb{F}_q (since \mathbf{H} is random), hence it is equal to \mathbf{s} with probability $q^{-(n-k)}$. Considering that $S_{\gamma,t}$ contains $\binom{n}{t}(2\gamma)^t$ vectors, and that at least one solution always exists by hypothesis, we obtain the result in (1). For the maximum support size version of R-SDP (i.e., R-SDP $_{\gamma,\leq t}$), we instead consider $\mathbf{s} \stackrel{\$}{\leftarrow} \{\mathbf{e}\mathbf{H}^\top \mid \mathbf{e} \in B_{\gamma,t}\}$, and consequently estimate the number of solutions as

$$N_{\gamma,\leq t} = 1 + \frac{|B_{\gamma,t}| - 1}{q^{n-k}} \approx 1 + \sum_{i=0}^t \binom{n}{i} 2^{i(1+\log_2(\gamma)) - (n-k)\log_2(q)}. \quad (2)$$

We will furthermore distinguish between two cases, depending on the relation between t and n . In the so-called *small support case* we have $t \ll n$, while in the *large support case* we have that t is close to n .

5.1 Solving the R-SDP with fixed and small support

When the support of the searched vector is small, one may solve the R-SDP through approaches following the principle of Information Set Decoding (ISD). This family of techniques, which are the best general solvers for the SDP in the Hamming metric, exploit the presence of a large number of null coordinates in the searched vector. Prange’s algorithm [28] historically dates as the first ever proposed ISD algorithm and can be used to decode both binary and non-binary codes with essentially the same complexity (if we neglect the cost of linear algebra in a non binary finite field). During the years, many improved ISD algorithms have been proposed, aimed at reducing the computational cost (see [4] for an overview of ISD algorithms for the binary field). Roughly, the main idea is that of increasing the guessing probability by allowing the presence of some non-null entries in the information set, and then proceed to identify them through enumeration. Combining this idea with collision search techniques, the complexity of ISD can be significantly reduced (see [8] and [26, 20] for state-of-the-art ISD algorithms for the binary and non-binary cases, respectively).

Let us now consider the R-SDP with one solution having support size $t \ll n$. First, Prange’s algorithm can be used without any modification, since it only searches for a sufficiently large number of zero positions. However, we can do better by relying on more advanced ISD algorithms. In fact, a solution for the R-SDP is also a solution for the corresponding SDP instance. Yet, one can probably optimize the algorithm by taking into account some observations:

1. when enumerating the candidates one should take into account that the vector is restricted and hence takes values in $\{0, \dots, \pm\gamma\}$. This leads to a polynomial reduction in the ISD complexity, with respect to the general case in the Hamming metric;
2. when the set entries of a candidate vector are not equally distributed over $\{\pm 1, \dots, \pm\gamma\}$, then we can further speed-up the enumeration phase. Indeed, if we know that some values are very unlikely to appear, then we can remove them from the search, achieving another polynomial advantage in the cost.

Based on the above considerations, the complexity of solving R-SDP $_{\gamma,=t}$ with small support can be bounded between the cost of binary ISD algorithms (as lower bound) and that of non-binary ISD algorithms (as upper bound) for the same support size. Indeed, the easiest R-SDP instance is the one in which the set entries of the solution have all the same value (say, are all equal to 1). In such a case, the problem is identical to the binary SDP, with the only exception that the considered code lives in a non-binary finite field (hence, the cost of linear algebra is slightly larger). Employing the well-known approximation for the cost of a binary ISD due to Canto-Torres and Sendrier [32], we can then conservatively

assume that solving R-SDP with small support costs at least

$$C_{ISD}(n, k, t) = 2^{-t \log_2(1-k/n) - \log_2(N_{\gamma,=t})}, \quad (3)$$

where the term $\log_2(N_{\gamma,=t})$ takes into account the existence of multiple solutions. Notice that, if $t \ll n$, then $N_{\gamma,=t} \approx 1$, so that the reduction in the complexity does not take place.

5.2 Solving the R-SDP with large support

When the support size of the solution is rather large (say, close to n), ISD becomes ineffective since the searched vector does not contain a large number of null entries. We here generalize the approach of [5], which solves the R-SDP with fixed and maximum support, for the sole case of $\gamma = 1$.

To this end, to solve instances of R-SDP $_{\gamma, \leq t}$ with t being close to n , we propose to use Algorithm 1, which is a natural generalization of the one proposed in [5]. Basically, the algorithm first brings the given parity-check matrix into par-

Input: $\mathbf{H} \in \mathbb{F}_q^{(n-k) \times n}$, $\mathbf{s} \in \mathbb{F}_q^{n-k}$, $\ell \in [0; n-k]$
Output: $\mathbf{e} \in B_{\gamma, t}$ such that $\mathbf{e}\mathbf{H}^\top = \mathbf{s}$

- 1 Pick a random permutation π .
- 2 Find \mathbf{A} such that $\mathbf{A}\pi(\mathbf{H}) = \begin{bmatrix} \mathbf{I}_{n-k-\ell} & \mathbf{H}' \in \mathbb{F}_q^{(n-k-\ell) \times (k+\ell)} \\ \mathbf{0}_{\ell \times (n-k-\ell)} & \mathbf{H}'' \in \mathbb{F}_q^{\ell \times (k+\ell)} \end{bmatrix}$; if it is not possible, restart from line 1.
- 3 Compute $[\mathbf{s}', \mathbf{s}''] = \mathbf{A}\mathbf{s}$, with $\mathbf{s}' \in \mathbb{F}_q^{n-k-\ell}$ and $\mathbf{s}'' \in \mathbb{F}_q^\ell$.
- 4 Produce a set $E_{\gamma, \ell} \in S_{\gamma, k+\ell}$ of solutions to the R-SDP $_{\gamma, k+\ell}$ instance represented by $\{\mathbf{s}'', \mathbf{H}''\}$.
- 5 **for** $\mathbf{e}'' \in E_{\gamma, \ell}$ **do**
- 6 Compute $\mathbf{e}' = \mathbf{s}' - \mathbf{e}''\mathbf{H}''^\top$
- 7 **if** $\mathbf{e}' \in S_{\gamma, t-k-\ell}$ **then**
- 8 **return** $\pi^{-1}([\mathbf{e}', \mathbf{e}''])$
- 9 Restart from line 1.

Algorithm 1: PGE+SS approach to solve R-SDP $_{\gamma, t}$

tially row-reduced echelon form, via a column permutation π and row operations described by the full-rank matrix \mathbf{A} . The same transformation is applied to the syndrome \mathbf{s} (line 3 of the Algorithm), in order to obtain a length- ℓ sub-syndrome \mathbf{s}'' which, together with \mathbf{H}'' , is given as input to a R-SDP $_{\gamma, k+\ell}$ solver (line 4 of the Algorithm). Finally, the found solutions, which are grouped in a set $E_{\gamma, \ell}$, are tested aiming to produce a solution to the initial R-SDP $_{\gamma, =t}$ instance (lines 5–8 in the Algorithm).

Proposition 2. Let $\{\mathbf{H}, \mathbf{s}\} \in \mathbb{F}_q^{(n-k) \times n} \times \mathbb{F}_q^{n-k}$ be an $R\text{-SDP}_{\gamma, \leq t}$ instance, with \mathbf{H} being a parity-check matrix of a random code $\mathcal{C} \subseteq \mathbb{F}_q^n$ with length n and dimension k , and $\mathbf{s} \stackrel{\$}{\leftarrow} \{\mathbf{eH}^\top \mid \mathbf{e} \in B_{\gamma, t}\}$. Then, Algorithm 1 solves $R\text{-SDP}_{\gamma, =t}$ with an average cost of

$$O \left(\frac{T(t, \gamma, \ell) + \frac{M(t, \gamma, \ell)}{1 + \widetilde{M}(t, \gamma, \ell)}}{\eta(t, \gamma, \ell) \left(1 - \prod_{i=k+\ell}^t \left(1 - \binom{i}{k+\ell} / \binom{n}{k+\ell} \right)^{N_{\gamma, =i}} \right)} \right),$$

where $N_{\gamma, =i}$ and $N_{\gamma, \leq t}$ are as in (1), and (2), $T(t, \gamma, \ell)$ is the average cost of an algorithm that produces $M(t, \gamma, \ell)$ solutions to an instance of $R\text{-SDP}_{\gamma, =k+\ell}$, and $\widetilde{M}(t, \gamma, \ell)$ out of these solutions lead to a success of Algorithm 1. Finally, $\eta(t, \gamma, \ell)$ denotes the probability that $\widetilde{M}(t, \gamma, \ell)$ is not null.

To conclude the analysis, we have to consider the cost of solving the small $R\text{-SDP}_{\gamma, k+\ell}$ instance represented by $\{\mathbf{H}'', \mathbf{s}''\}$. As in [5], we consider the application of Wagner's algorithm [34], originally proposed as a solver for the subset sum problem. Among the solutions to the problem, we assume that

1. the number of *good solutions*, i.e., vectors leading to a success for Algorithm 1, is given by

$$U_{\gamma, t, \ell} = \sum_{i=k+\ell}^t \frac{\binom{i}{k+\ell}}{\binom{n}{k+\ell}} N_{\gamma, =i}; \quad (4)$$

2. the number of *bad solutions*, i.e. vectors that do not lead to a success of Algorithm 1, is given by

$$U'_{\gamma, y, \ell} = \max \{0, q^{-\ell} ((2\gamma)^{k+\ell} - U_{\gamma, t, \ell})\}. \quad (5)$$

Wagner's algorithm on a levels to solve the $R\text{-SDP}$ with maximum weight associated to $\{\mathbf{H}'', \mathbf{s}''\}$ is detailed in Algorithm 2. For the sake of simplicity, we assume that $k+\ell$ is a multiple of 2^a ; \mathbf{H}''_i , for $i \in [0; 2^a - 1]$, denotes the matrix formed by the columns of \mathbf{H}'' in the positions $\{i \frac{k+\ell}{2^a}, \dots, (i+1) \frac{k+\ell}{2^a} - 1\}$. In the algorithm, the merging operation between two lists, which we denote as $\mathcal{L}_{2j}^{(i)} \sqcap_{u_i} \mathcal{L}_{2j+1}^{(i)}$, is defined as follows

$$\{(\mathbf{z}_{2j} + \mathbf{z}_{2j+1}, [\mathbf{p}_{2j}, \mathbf{p}_{2j+1}]) \mid (\mathbf{z}_b, \mathbf{p}_b) \in \mathcal{L}_b^{(i)}, \mathbf{z}_{2j} + \mathbf{z}_{2j+1} = \mathbf{0} \text{ in the last } u_i \text{ entries}\}.$$

Similarly to [5, Proposition 16], in the following proposition we assess the complexity of using Wagner's algorithm to find one of the desired solutions.

Proposition 3. Let $\{\mathbf{H}'', \mathbf{s}''\} \in \mathbb{F}_q^{\ell \times (k+\ell)} \times \mathbb{F}_q^\ell$ be an $R\text{-SDP}_{\gamma, =k+\ell}$ instance with $U_{\gamma, t, \ell}$ good solutions and $U'_{\gamma, t, \ell}$ bad solutions. Assume to run Algorithm 2 on a levels, with options $0 = u_0 < u_1 < \dots < u_{a-1} < u_a = \ell$.

<p>Input: $\mathbf{H}''_0, \dots, \mathbf{H}''_{2^a-1} \in \mathbb{F}_q^{\ell \times \frac{k+\ell}{2^a}}$, $\mathbf{s}'' \in \mathbb{F}_q^\ell$</p> <p>Output: A list $\mathcal{L}_0^{(a)} = \{(\mathbf{p}\mathbf{H}''^\top, \mathbf{p})\}$ such that $\mathbf{p} \in \{\pm 1, \dots, \pm \gamma\}^{k+\ell}$ and $\mathbf{p}\mathbf{H}''^\top = \mathbf{s}''$</p> <p>Data: $v, a \in \mathbb{N}$, with $a \geq 1$ and $v \leq \frac{k+\ell}{2^a}$, and positive integers $0 < u_1 < \dots < u_{a-1} < \ell$.</p> <ol style="list-style-type: none"> 1 Set $u_0 = -1, u_a = \ell$. 2 Choose random subsets $\mathcal{R}_0, \dots, \mathcal{R}_{2^a-1} \subseteq \{\pm 1, \dots, \pm \gamma\}^{(k+\ell)/2^a}$, each of size $(2\gamma)^v$. 3 Build the lists $\mathcal{L}_j^{(0)} = \{(\mathbf{z} = \mathbf{p}\mathbf{H}_j''^\top, \mathbf{p}) \mid \mathbf{p} \in \mathcal{R}_j\}$ for $j \in [0; 2^a - 2]$. 4 Build the list $\mathcal{L}_{2^a-1}^{(0)} = \{(\mathbf{z} = \mathbf{p}\mathbf{H}_{2^a-1}''^\top - \mathbf{s}'', \mathbf{p}) \mid \mathbf{p} \in \mathcal{R}_{2^a-1}\}$. 5 for $i = 1$ to a do 6 for $j = 0$ to $2^{a-i} - 2$ do 7 $\mathcal{L}_j^{(i+1)} = \mathcal{L}_{2j}^{(i)} \sqcap_{u_i} \mathcal{L}_{2j+1}^{(i)}$ 8 return $\mathcal{L}_0^{(a)}$
--

Algorithm 2: Wagner's algorithm structured on a levels

Let $\rho = 2^{(2^a v - k - \ell - \log 2(q)) (1 + \log_2(\gamma)) - \log_2(q) \sum_{i=1}^{a-1} u_i 2^{a-1-i}}$. Then, the computational complexity used by Wagner's algorithm is given by

$$T(k, \ell, \gamma) = \min_{\substack{a, v, u_1, \dots, u_{a-1} \in \mathbb{N} \\ a \geq 1, v \leq \frac{k+\ell}{2^a} \\ 0 < u_1 < \dots < u_{a-1} < \ell}} \left\{ \max_{i \in [0; a]} \left\{ 2^{v 2^i \log_2(\gamma) - \phi(i) \log_2(q)} \right\} \right\},$$

where $\phi(i) = \begin{cases} 0 & \text{if } i = 0, \\ u_i + \sum_{m=1}^{i-1} 2^{i-1-m} u_m & \text{otherwise.} \end{cases}$

The algorithm finds good solutions with probability $\eta(t, \gamma, \ell) = 1 - (1 - \rho)^{U_{\gamma, t, \ell}}$, and on average outputs $M(t, \gamma, \ell) = \rho \left(U_{\gamma, t, \ell} + U'_{\gamma, t, \ell} \right)$ solutions, with $\widetilde{M}(t, \gamma, \ell) = \rho U_{\gamma, t, \ell}$ of them being good.

6 Security and practical instances

Due to lack of space, we do not provide a formal proof of security, but simply hints at how such a proof should work, by highlighting possible attack strategies. The analysis of such attack avenues also allows us to describe how the system parameters can be designed.

One-wayness of key generation The public key is a collection of syndromes of restricted vectors. Indeed, each row of the public \mathbf{S} is the syndrome of the corresponding row of the secret \mathbf{E} . Notice that, in the key generation algorithm, we guarantee that the rows have minimum support size in the range t_E , with $t_E \ll n$. Since the entries of \mathbf{E} are either null or equal to ± 1 , we have that

finding each row of \mathbf{E} can be seen as facing an R-SDP $_{1, \leq t^*}$ instance, for some $t^* \geq t_E$. As we have highlighted in Section 5, the computational complexity to solve R-SDP in case of a small support, for a code \mathcal{C} , is not lower than that of solving the Hamming SDP for a code with same length and dimension of \mathcal{C} , but defined over the binary finite field, searching for a vector with Hamming weight t^* . As (3) shows, the complexity grows exponentially with the weight of the searched vector. Hence, conservatively, we assess the complexity of attacks aimed at recovering the secret key as $2^{-t_E \log_2(r/n)} - \log_2(N_{1, t_E})$.

Unforgeability To forge a signature, an attacker may proceed as follows. First, he picks a random \mathbf{y} with the desired distribution and computes $\mathbf{s}_y = \mathbf{y}\mathbf{H}^\top$. Then, he sets $\mathbf{c} = \text{Hash}(m, \mathbf{s}_y)$ and computes $\mathbf{s}_z = \mathbf{s}_y + \mathbf{c}\mathbf{S}$. If he/she is able to produce a vector $\mathbf{z} \in B_{\gamma, n}$ such that $\mathbf{z}\mathbf{H}^\top = \mathbf{s}_z$, then the pair $\{\mathbf{c}, \mathbf{z}\}$ can be used as a valid signature. Notice that, to do this, he/she must solve an R-SDP $_{\bar{\gamma}, \leq n}$ instance. To assess the hardness of this attack, we rely on Proposition 2.

Unfeasibility of noise recovery Assume that the adversary is able to retrieve \mathbf{y} from \mathbf{s}_y . If he succeeds, he/she can then compute $\mathbf{z} - \mathbf{y} = \mathbf{c}\mathbf{E}$ and, since \mathbf{c} is known, retrieve information on \mathbf{E} . Exploiting the sparsity of both \mathbf{c} and \mathbf{E} , the rows of \mathbf{E} can be trivially recovered. Then, we have to guarantee that recovering \mathbf{y} is unfeasible. Again, this reduces to the problem of solving an R-SDP $_{\gamma, \leq n}$ instance (hence, we use Proposition 2 to estimate the complexity of attacks of this kind).

Statistical indistinguishability Finally, we consider the possibility for an attacker of retrieving some information about the secret key by performing a statistical analysis on a bunch of collected honest signatures. This is motivated by the fact that the same private key is used to construct many signatures, and each signature $\{\mathbf{c}, \mathbf{z}\}$ is somehow related to the secret \mathbf{E} . To describe how such a dependence can be exploited, consider a collection of signatures for which the digests \mathbf{c} have a common set entry. To analyze this situation, we assume for simplicity that such an entry is the first one, and that $\bar{\gamma} = \gamma$. In such a case, the first row of \mathbf{E} contributes to all the collected signatures, and its entries can be recovered through a statistical analysis. Indeed, in the positions $i \in [0; n - 1]$ such that $e_{0, i} = \pm 1$, we have that the i -th entry of the signatures takes value equal to $\pm \bar{\gamma}$ with a slightly lower probability than $(2\bar{\gamma} + 1)^{-1}$. If $e_{0, i} = 0$, on the contrary, this statistical bias is not present. Hence, collecting a sufficient number of signatures with such digests would be enough to recover the first row of \mathbf{E} . Thanks to the rejection sampling in the signing algorithm, it is enough to choose $\bar{\gamma} \leq \gamma - 1$ to prevent this attack. Yet, this attack can be generalized by considering the occurrence, in the digests, of specific tuples of size larger than $\gamma - \bar{\gamma}$. Indeed, a pattern of this size may be such that its product with a column of \mathbf{E} yields a value larger than $\gamma - \bar{\gamma}$ (or lower than $-(\gamma - \bar{\gamma})$): in such a case, the statistical bias in the signatures appears, and it can somehow be used to recover information about the secret \mathbf{E} . To completely prevent from this kind

of attacks, it is enough to choose $\gamma - \bar{\gamma} \geq \min\{w_c, w_E\}$, which is the maximum absolute value of each entry of \mathbf{cE} .

6.1 Parameter choices

In order to design secure parameters for the new system, we must first guarantee that the number of possible digests is sufficiently large, that is, $\binom{b}{w_c} 2^{w_c} > 2^\lambda$. Then, in order to choose w_E , we first consider (3) and derive the minimum value of t_E guaranteeing that the complexity of ISD is higher than 2^λ . The number of times the key generation algorithm has to be repeated on average, before obtaining a matrix \mathbf{E} where each row has support size at least t_E , can be estimated with simple combinatorial arguments, and is given by $\left(1 - \sum_{i=0}^{t_E-1} \binom{n}{i} \left(\frac{w_E}{b}\right)^i \left(1 - \frac{w_E}{b}\right)^{n-i}\right)^{-1}$. Taking all of this into account, in Table 1 we provide some parameter sets achieving $\lambda = 128$ bits of classical security.

As we see from the table, the scheme achieves very compact signatures and public keys, and also enables flexible trade-offs between the public key size and the computational efficiency (which is inversely proportional to the average number of rejected signatures per each generated signature).

(n, k, q)	b	w_E	w_c	t_E	γ	$\bar{\gamma}$	Avg rejections	$ \sigma $ in kB	$ \text{pk} $ in kB
(700, 450, 16381)	150	24	28	87	4300	4272	96.79	1.09	65.62
(900, 750, 16381)	150	12	28	50	5500	5472	98.8	1.43	39.37
(900, 450, 16381)	5500	26	28	128	5500	5472	98.8	1.43	118.12
(900, 700, 16381)	100	9	37	59	5400	5363	486.54	1.44	35.00
(900, 550, 16381)	200	26	25	94	6000	5974	49.81	1.45	122.50

Table 1. Instances achieving 128 bits of classical security. For all the considered parameters sets, the rejection rate in the key generation algorithm is lower than 0.1, implying that on average the algorithm has to be repeated for no more than 10 times.

7 Conclusion

We have proposed a novel adaptation of the Schnorr-Lyubashevsky approach to the design of digital signature schemes based on codes in the Hamming metric. By relying on large weight vectors with restricted entries, the proposed scheme is able to withstand known cryptanalysis approaches, while achieving very compact signatures and keys. In this first proposal, we have considered random-like non-structured codes, which allow relying on the general formulation of the corresponding decoding problems for the security of the scheme. A formal security proof, along with the study of variants adopting structured codes (e.g., quasi-cyclic codes), which may achieve further reductions in the public key size, is left for future works.

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Appendix A - Proofs of Section 4

Proof of Lemma 1

Proof. The probability is the same for β and $q - \beta$, so we only consider the case of $\beta \in [0; \lfloor q/2 \rfloor]$, for which $\beta_q = \beta$. Let u_1 denote the number of indexes i for which $c_i e_i = 1$, and u_{-1} denote that of indexes for which $c_i e_i = -1$. To have $\langle \mathbf{c}; \mathbf{e} \rangle = \beta$, it must $u_1 - u_{-1} = \beta$. Let $v = u_1 + u_{-1}$, that is, the number of intersections between the support of \mathbf{c} and that of \mathbf{e} . To have $\langle \mathbf{c}; \mathbf{e} \rangle = \beta$, the following two conditions must be verified:

- we have $\max\{0, w_c + w_E - \ell\} \leq v \leq \min\{w_E, w_c\}$;
- since $u_{-1} = v - u_1$, it must be $u_1 = \frac{v+\beta}{2}$.

Hence, we obtain the following probability:

$$\Pr[\langle \mathbf{a}; \mathbf{b} \rangle = \beta] = \sum_{\substack{v=\beta \\ v \text{ and } \beta \text{ have the same parity}}}^{\min\{w_E, w_c\}} 2^{-v} \frac{\binom{v}{\frac{v+\beta}{2}} \binom{w_E}{v} \binom{\ell-w_E}{w_c-v}}{\binom{\ell}{w_c}}.$$

□

Proof of Lemma 2

Proof. The proof is straightforward. Indeed, assume that $y = x$, which happens with probability $1/(2\alpha + 1)$. To have $\langle \mathbf{c}; \mathbf{e} \rangle + x = \beta$, it must be $\langle \mathbf{c}; \mathbf{e} \rangle = \beta - x$. Summing over all possible values of x , we obtain the formula in the thesis. □

Appendix B - Proofs of Section 5

Proof of Proposition 2

For a vector $\mathbf{e} \in S_{\gamma, i}$ for which $\mathbf{e}\mathbf{H}^\top = \mathbf{s}$, the probability that π is such that the vector \mathbf{e}'' formed by the last $k + \ell$ entries of $\pi(\mathbf{e})$ are all non null is given by $\epsilon_i = \binom{i}{k+\ell} / \binom{n}{k+\ell}$. Notice that, if \mathbf{e}'' has a different support size, then it will never be in the set $E_{\gamma, \ell}$ produced in Line 4 and, hence, we have that \mathbf{e} is never returned as output from Algorithm 1. The number of solutions with support equal to i can be estimated as $N_{\gamma, =i}$, so that the probability that π is not valid for all of them is obtained as $(1 - \epsilon_i)^{N_{\gamma, =i}}$. Multiplying over the values of i from $k + \ell$ to t , and taking the complementary, we derive the probability that π is valid for at least one out of the $N_{\gamma, \leq t}$ solutions. We multiply this probability by $\eta(t, \gamma, \ell)$ and obtain the success probability of one iteration of Algorithm 1.

Let $T(t, \gamma, \ell)$ denote the cost of an R-SDP $_{\gamma,=k+\ell}$ solver (i.e., for the instance represented by $\{\mathbf{H}'', \mathbf{s}''\}$), producing a set of solutions $E_{\gamma, \ell}$ that on average contains $M(t, \gamma, \ell)$ elements. Since there are $\widetilde{M}(t, \gamma, \ell)$ actually valid solutions (i.e., leading to the success of Algorithm 1), on average we test $M(t, \gamma, \ell) / (1 + \widetilde{M}(t, \gamma, \ell))$ vectors from $E_{\gamma, \ell}$, before Algorithm 1 successfully halts. Thus, we estimate the cost of executing lines 4–8 as $T(t, \gamma, \ell) + M(t, \gamma, \ell) / (1 + \widetilde{M}(t, \gamma, \ell))$.

Proof of Proposition 3

Let $\mathbf{x} \in \{\pm 1, \dots, \pm \gamma\}^{k+\ell}$ be a solution, i.e., such that $\mathbf{x}\mathbf{H}''^\top = \mathbf{s}''$. Wagner's algorithm will output \mathbf{x} among the vectors in the final list iff

- a) $\mathbf{x} \in \mathcal{R}_0 \times \mathcal{R}_1 \times \dots \times \mathcal{R}_{2^a-1}$;
- b) at each merge, the vector \mathbf{x} is not filtered.

Condition a) is verified with probability $\left(\frac{(2\gamma)^v}{(2\gamma)^{\frac{k+\ell}{2^a}}}\right)^{2^a} = 2^{(2^a v - k - \ell)(1 + \log_2(\gamma))}$.

We now proceed by computing the probability that also condition b) happens, assuming condition a) holds. If $a = 1$, then we have no filtering, while in the other cases it may happen in the levels from the first to the $(a - 1)$ -th one. To this end, we consider the i -th level (for $i \in [1; a]$), and divide \mathbf{x} into 2^{a-i+1} chunks, each formed by $\frac{k+\ell}{2^{a-i+1}}$ consecutive entries, which we denote as \mathbf{x}_j , for $j \in [0; 2^{a-i} - 1]$. In the i -th level, \mathbf{x} will not be filtered if and only if

- 1. for $j \in [0; 2^{a-i} - 2]$, $\mathbf{x}_{2j}\mathbf{H}''_{2j}^\top + \mathbf{x}_{2j+1}\mathbf{H}''_{2j+1}^\top$ is null in the last $u_i - u_{i-1}$ positions;
- 2. $\mathbf{x}_{2^{a-i+1}-2}\mathbf{H}''_{2^{a-i+1}-2}^\top + \mathbf{x}_{2^{a-i+1}-1}\mathbf{H}''_{2^{a-i+1}-1}^\top - \mathbf{s}''$ is null in the last $u_i - u_{i-1}$ positions.

Note that if condition 1 is met, then condition 2 is met as well, so we just have to consider the probability with which condition 1 happens. Given that both \mathbf{H}'' and \mathbf{x} are random, in each merge, chunks \mathbf{x}_{2j} and \mathbf{x}_{2j+1} will not be filtered out with probability $q^{-(u_i - u_{i-1})}$. Given that, for $j \in [0; 2^{a-i} - 2]$, we perform $2^{a-i} - 1$ merges, condition 1 is verified with probability

$$\left(q^{-(u_i - u_{i-1})}\right)^{2^{a-i} - 1} = 2^{-(2^{a-i} - 1)(u_i - u_{i-1}) \log_2(q)}.$$

Thus, the probability of \mathbf{x} surviving till the last level is obtained as

$$\prod_{i=1}^{a-1} \left(q^{-(u_i - u_{i-1})}\right)^{2^{a-i} - 1} = 2^{-\log_2(q) \sum_{i=1}^{a-1} u_i 2^{a-1-i}},$$

where we consider $u_0 = 0$. Putting everything together, we get that each solution \mathbf{x} is found with probability

$$\rho = 2^{(2^a v - k - \ell)(1 + \log_2(q)) - \log_2(q) \sum_{i=1}^{a-1} u_i 2^{a-1-i}}.$$

The total number of produced solutions is then estimated as

$$M(k, \ell, \gamma) = \rho(U_{\gamma, t, \ell} + U'_{\gamma, t, \ell}),$$

while the success probability is simply obtained by considering the probability that at least one good solution survives till the end, that is

$$\eta(\gamma, t, \ell) = 1 - (1 - \rho)^{U_{\gamma, t, \ell}}.$$

We finally derive the computation complexity to execute Wagner's algorithm. To this end, we assume that the cost of each merge is equal to the list size. In the initial level, we use lists of size $L_0 = (2\gamma)^v$. In the first level, i.e., for $i = 1$, the average size of the lists is given by $L_1 = L_0^2/q^{u_1} = 2^{2v(1+\log_2(\gamma)) - u_1 \log_2(q)}$. In the i -th level, for $i \geq 1$, the average size of the lists $\mathcal{L}_j^{(i)}$ is $L_i = \frac{L_{i-1}^2}{q^{u_i - u_{i-1}}}$. With simple computations, we find that

$$L_i = 2^{v2^i(1+\log_2(\gamma)) - \phi(i) \log_2(q)},$$

with $\phi(i) = \begin{cases} 0 & \text{if } i = 0, \\ u_i + \sum_{m=1}^{i-1} 2^{i-1-m} u_m & \text{otherwise.} \end{cases}$

We use the maximum of these quantities as a conservative lower bound on the complexity.